APPLICATION OF BOX-JENKINS MODELS TO THE TOURIST INFLOW IN BHUTAN

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Abstract

Bhutan has now increasingly become a popular destination for many international tourists. Tourism in Bhutan is considered as one of the largest foreign earning industries. The number of tourist inflow in the country is increasing year by year. Forecasting is very necessary for administration and tourist agent for creating awareness and planning for the future development. It can also predict the future trends as accurately as possible and helps in staying one step ahead of the competition. This study aims to apply mathematical model for forecasting monthly tourist inflow from Malaysia, Singapore, China, USA, England, France, Germany, Thailand, Australia and Japan to Bhutan. The Box-Jenkins model is used to identify the parameters of Autoregressive integrated moving average (ARIMA) model of monthly tourist visited data of above mentioned countries in the period 2011-2015 obtained from Tourism Council of Bhutan. An Akaike's Information Criterion, Schwartz's Bayesian Criterion and estimate variance of white noise are used throughout to test for the identification of best fit model. Further, the periodogram analysis was used to confirm the seasonal period of the model. The results showed ARIMA model for Thai, Chinese, Malaysian and Japanese, while seasonal ARIMA for American, Australian, British, French, Singaporean and German. Further, seasonal ARIMA model was obtained as the best fit model for the overall data. These models are illustrated and could possibly forecast the monthly tourist inflow of one year ahead with acceptable accuracy.

Keywords: Akaike's information criterion, Box-Jenkins model, Schwartz's Bayesian criterion, variance of white noise

1. Introduction

Bhutan's tourism industry began in 1974. It was introduced with the primary objective of generating revenue, especially foreign exchange; publicizing the country's unique culture and traditions to the outside world, and to contribute to the country's socio-economic development. Bhutan has now increasingly become a popular destination for many international tourists (Dorji, 2014). The number of tourist inflow in the country is increasing year by year. So forecasting is

very necessary for administration and tourist agent for creating awareness and planning for the future development (Honey & Gilpin, 2009). It can also predict the future trends as accurately as possible and helps in staying one step ahead of the competition (Singha, 2012).

In spite of the great importance of tourism in Bhutan's economy, there is an incomprehensible lack of systematic and up to date quantitative research oriented on analyzing the core determinant and patterns of the visitors from different countries in Bhutan. Therefore it is very important to analyze the determinant and the core pattern tourist inflow in Bhutan. Such a study can be used in formulation of future macroeconomic development strategies, pricing strategies and tourism routing strategies in Bhutan as a one of the most popular destination in the world (Mahmood & Ali, 2016).

Given the importance of the tourism industry to Bhutan it is essential to generate the accurate forecast of the future trends of tourist flows from the major origin countries. This research paper aimed is one-period-ahead forecasts of international tourism demand for Bhutan, and to seek provide the best model for forecasting international tourist arrivals to Bhutan for these periods using Box Jenkins Methods with non-seasonal and seasonal modification. The advantages of Box Jenkins Methodology involve selecting a great quantum of information from the analyzed empirical time series, using a small number of parameters (Nanthakumar & Ibrahim, 2010).

2. Data and methodology

2.1 Data

For the study, we used the monthly tourist inflow series includes data from top ten market sources that is Malaysia, Singapore, China, America, Australia, USA, France, Thailand, Germany and Japan during 2011-2015 obtained from Tourism Council of Bhutan. (Tourism Council of Bhutan, n.d) The series consist of 60 observations where first 54 sample observations are analyzed by using the Box-Jenkins method and the rest 6 observations out of sample are used to compare the forecast in all the nationalities.

2.2 Methodology

In this study, the Box-Jenkins model is used to analyze tourist inflow series which include data from top ten market sources that is Malaysia, Singapore, China, America, Australia, USA, France, Thailand, Germany, Japan and total tourist visitedduring 2011 – 2015. The Box-Jenkins method consists of following steps. (Singh, 2013).

The first step is to identify the all tentative model. Identification consists of specifying appropriate AR, MA or ARMA and order of the model. The identification is done by looking at the ACF and PACF of the interested stationary series.

The second step is to estimate the parameters of the model. Parameters of models can be estimated by least-square method. The estimation of parameter usually requires more complicated iteration procedure but the computer programming automatically generate it.

The third step is to check the model. This step is also called diagnostic checking or verification. Two important elements in checking are to ensure that the residuals of the model are random and white noise i.e. uncorrelated and constant variance $\hat{\sigma}_a^2$, and also to make sure that the estimated parameters are strictly significant (Newbold, 2013).

The fourth step is to elect the best model from the various ARIMA models which might be suitable for the series. Thus we use Akaike's Information Criteria (AIC) and Schwarz's Bayesian Criterion (SBC) for model selection to find the best model of the monthly tourist inflow in Bhutan.

2.2.1 ARIMA model

If the process is not stationary, we have to take differencing term $(1-B)^d$ in the process. When model ARMA (p, q) model on a time series which has been differenced d times we call this an ARIMA (p, d, q) model [4]. Thus the general ARIMA (p, d, q) written using backshift operator

$$f_{p}(B)(1-B)^{d} X_{t} = q_{0} + q_{q}(B)a_{t}$$

where $\phi_p(B) = 1 - \phi_1 B - \phi_2 B^2 - \dots - \phi_p B^p$ is the stationary AR operator,

as:

$$\theta_q(B) = 1 - \theta_1 B - \theta_2 B^2 - \dots - \theta_q B^q$$
 is the invertible MA operator

and $\theta_0 = \mu(1 - \phi_1 - \phi_2 - \dots - \phi_p)$ which is known as deterministic trend term. (Ekpenyong, 2016)

2.2.2 Seasonal ARIMA

The Seasonal ARIMA model includes both seasonal and non-seasonal factors in multiplicative model called ARIMA $(p, d, q) \times (P, Q, D)_s$ where p is non-seasonal AR order, d is non-seasonal differencing, q is non-seasonal MA order, P is seasonal AR order, P is seasonal differencing, Q is seasonal MA orderand P is seasonal periodwhere the general equation of the model is

$$\Phi_P(B^s)\phi_P(B)(1-B)^d(1-B^s)^DX_t = \theta_q(B)\Theta_Q(B)^sa_t$$

(Baldigara & Mumula, 2015) where $\phi_p(B)$ is the non-seasonal autoregression component of order p, $\Phi_p(B^s)$ is the seasonal autoregression of order s, X_t is the current value of the time series examined, B is the backward shift operator $X_t(B^i) = X_{t-i}$, $(1-B)^d$ is non-seasonal difference term, $(1-B^s)^D$ is the seasonal difference term, $\theta_q(B)$ is the non-seasonal moving average of order q and $\Theta_Q(B)^s$ is the seasonal moving average of order Q (1990, p. 106)

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2.2.3 Periodogram analysis

To confirm the seasonality, we perform periodogram analysis. Periodogram help us to find the hidden periodicities. If model contains a single periodic component at frequency ω , the periodogram $I(\omega_k)$ at Fourier frequency ω_k closest to ω will be maximum. Thus the maximum periodogram ordinate will be

$$I^{(1)}(\omega_{(1)}) = \max \left\{ I(\omega_k) \right\}$$

where $\omega_{(1)}$ is to indicate the maximum periodogram ordinate of the Fourier frequency.

Under the null hypothesis H_0 , including a period in a time series, an exact test statistic of $I^{(1)}(\omega_{(1)})$ is known as Fisher's test which is based on the equation below

$$T = \frac{I^{(1)}(\omega_{(1)})}{\sum_{k=1}^{[n\ 2]} I(\omega_k)}.$$

3. Results

The monthly tourist visited from top ten market source during 2011 – 2015 is analyzed by using Box – Jenkins's model. Out of which we will present only one nationality with all the process and rest will be presented with best fit models. The Box-Jenkins model consist of following steps: model identification, parameter estimation, diagnostic checking and model forecasting.

3.1.1 Model identification.

1. To use Box - Jenkins methodology, the series should be stationary. In this study, the graphical methods have been used to check the stationary of the series. In graphical method, graph and correlogram have been used. Figure 1 shows the graph of monthly American tourist visited

during 2011 – 2015. The data set suggests that the series is not stationary in variance. To stabilize the variance, we used Box-Cox transformation (Buthmann, n.d.). The preliminary residual mean square errors are calculated using the power transformation by SAS system software version 9.1 and we need logarithm transformation. Figure 2 shows the graph of monthly logarithm transformed American tourist visited which is stationary in the mean and variance. (Sample ACF and Properties of AR(1) Model. n.d.).

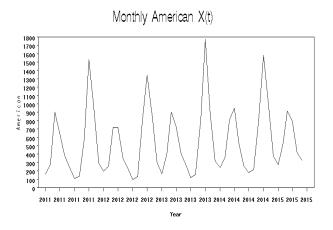


Fig. 1 The monthly American tourist visited during 2011 -2015

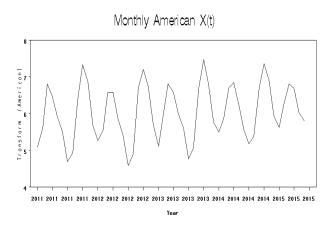


Fig. 2 The monthly logarithm transformed American tourist visited during 2011 -2015

Autocorre	latı	m

Partial Autocorrelations

Lag	Covariance	Correlation	-1 9 8 7 6 5 4 3 2 1 1	0 1 2 3 4 5 6 7 8 9 1	Std Error	Lag	Correlation	-1 9 8 7 6 5 4 3 2 1 0 1 2 3 4 5 6 7 8	8 9 1
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6	0.369330	0.65811	· ·	ste steateste steateste deste deste deste	0.245984	6 7	-0.11096	. **	!
7	0.166376	0.29647	1	stentedestesteste	0.276675		-0.04343	, " ,	!
8	-0.205447	36608	- Academic desirates		0.282496	8	-0.21393	1 1 1	
9	-0.200447	66030			0.291149	9	-0.08071	, ** .	- 1
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20	-0.169342	30175	· www.		0.407658	21	-0.05333	. *	- 1
21	-0.271148	48316	the desirable de		0.411774	22	-0.09942		- !
22	-0.117488	20935	, whitelest		0.422142			1 "" "	- !
23	0.158044	0.28162	1 .	*****	0.424060	23	0.04321	, j" ,	!
24	0.281238	0.50114	1 .	www.www.ww	0.427509	24	-0.10964	1 1 1 1	
25	0.116719	0.20620	i .	stratesteste	0.438263	25	-0.05546	. *1 .	
26	-0.134712	24004	. *****		0.440046	26	0.04109	. [*	
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28	-0.098576	17565	yestesteste		0.448629	28	0.00272	1 · 1 · ·	
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30	0.166823	0.29726	i .	stratestestesteste	0.450891	30	-0.04405	. * .	
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32	-0.119651	21321	violate tr	i . i	0.454908	32	0.00633	1	1
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34	-0.062858	11201	**	i . i	0.460501	34	0.00397	i i i	i
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36	0.136559	0.24333	i i	www.w	0.461900	36	0.03751	i . i* .	i i
37	0.050700	0.09034	i .	**	0.464267	37	-0.03745	*	i
38	-0.078288	13950	1 states	: 1	0.464593	38	-0.05589	* *1	- 1
39	-0.109674	19543	- Vertextext	: : : : : : : : : : : : : : : : : : : :	0.465368	39	0.04150		- 1
40	-0.038594	-,06877	*		0.466885	40	0.01359		- 1
41	0.044261	0.07887	<u> </u>		0.467073	41	-0.04034		
42	0.066537	0.11856	'		0.467319		0.02825	. " '	- !
42	0.000037	0.02646	1.1	' !	0.467876	42		. " .	!
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40	-0.000001	10039	1 . ""	• 1	0.400225	45	0.03890	. [* .	

Figure 3 Sample ACF and PACF of monthly

logarithm transformation of American tourist visited (Nau, n.d.).

Figure 3 shows sample ACF and PACF with 95% confidence limits. (Adhikari & Agrawal, 2013, pp. 1–3) The ACF shows damp sine cosine wave and slow decaying of the spikes indicates cyclic or seasonal movement of the correlation (Keshvani, 2013). Therefore to confirm the seasonality, we perform periodogram analysis. The periodogram analysis of the logarithm of monthly American tourist visited is clearly dominated by a very large peak at frequency, 1.04720. This frequency corresponds to a period of Pequals 6. It indicates that the data exhibit an approximate of 6 months cycle. So we need first seasonal differencing of period 6 months. Figure 4 shows sample ACF and PACF of monthly logarithm transformation of American tourist visited for first seasonal differencing at period 6 months. The sample ACF shows spike at lag 1, 6 and 12 and PACF cuts off after lag 6. Therefore, the tentative models are ARIMA $(1, 0, 0) \times (0, 1, 1)$ 6, ARIMA $(1, 0, 0) \times (0, 1, 2)$ 6, ARIMA $(4, 0, 0) \times (0, 1, 2)$ 6, ARIMA $(5, 0, 0) \times (0, 1, 2)$ 6 ARIMA $(0, 0, 1) \times (1, 1, 1)$ 6, ARIMA $(0, 0, 0) \times (1, 1, 0)$ 6 and ARIMA $(0, 0, 0) \times (0, 1, 1)$ 6.

			Autocorrelations				Partial Autocorrelat	ions
Lag	Covariance	Correlation	-1 9 8 7 6 5 4 3 2 1 0 1 2 3 4 5 6 7 8 9	1 Std Error	Lag	Correlation	-1 9 8 7 6 5 4 3 2 1	0 1 2 3 4 5 6 7 8 9 1
0	0.357732	1.00000	de dentente atentente dentente de dentente de	**1 0	1	0.56436	1	****
1	0.201891	0.56436	ate atentiente atentiente atentiente	0.144338	2	-0.42961	telefe telefetetetete	l · I
2	0.0092044	0.02573	j* ,	0.184674	3	0.30475	· 1	*****
3	-0.018725	05234	*	0.184749	4	-0.42808	telefe telefetetetetet	1
4	-0.039448	11027	**	0.185057	5	-0.53541	destrates tentes te	i , i
- 5	-0.182859	51116	the standard of the destate of the standard of	0.186421	6	-0.50208	de destente destente destente de	i i
6	-0.293355	82004	de desterte desterte desterte destertententente	0.213635	7	0.17642	i .	www.
7	-0.170188	47574	ste steriteite steriterbesterbeit	0.271403	8	-0.31660	statestatestate	i . i
8	-0.0097719	02732	. *	0.288253	9	-0.19094	teletele	i . i
9	0.0065530	0.01832	i .	0.288307	10	-0.05022	j . *	i . i
10	0.031425	0.08785	**	0.288331	11	-0.10420	, where	i . i
11	0.161116	0.45038	ste steakeste steakeste ste	0.288888	12	-0.02229	i .	i . i
12	0.250237	0.69951	de dedecte dedecte deste deste de	0.303164	13	-0.07101	· i · · · *	i . i
13	0.150897	0.42182	de dedede dededede	0.335107	14	-0.16395	***	i i i
14	0.0055975	0.01565		0.345992	15	0.02260	i	i i i
15	-0.0093078	-,02602	*	0.346006	16	-0.00795	i .	i . i
16	-0.020716	05791	*	0.346047	17	-0.09217	that.	i i
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19	-0.122262	34177	ate atenderate atende	0.375811	20	-0.17726	stestesteste	·
20	0.0010921	0.00305		0.382232	21	0.18563	-	'
21	0.011954	0.03342	. * .	0.382232	22	0.01350		·
22	0.020158	0.05635	. *	0.382293	23	-0.14269	totale	· '
23	0.109253	0.30541	the state of the s	0.382466	24	0.03465		' **
24	0.169620	0.47415	the absolute development of	0.387513	25	0.03964		' W
25	0.093110	0.26028	*******	0.399417	26	-0.02390		! ' !
26	0.00033299	0.00093		0.402935	27	-0.16242	· water	! ' !
27	-0.013350	03732	*	0.402935	28	0.08731		1
28	-0.025505	- , 07130	*	0.403007	29	0.00532		·
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31	-0.070402	-,19680	teletete ,	0.411562	32	-0.00999		! ' !
32	-0.0042451	01187		0.413518	33			 *
33	0.011790	0.03296	. *	0.413525		0.06146	totote	!" ' !
34	0.019678	0.05501	[*	0.413580	34 35	-0.12512		
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36	0.080204	0.22420	to the test of the second of t	0.414970	36	-0.05129	1 ***	! . !
37	0.046529	0.13007	the streets .	0.417486	37	-0.11818	. ""	
38	-0.0040485	01132		0.418330	38	-0.09584	!	
39	-0.012743	03562	*!	0.418336	39	-0.00343	·	!
40	-0.012317	03443	*	0.418399	40	-0.10230	. ***	
41	-0.034025	09511	***	0.418458	41	-0.13557	. ***	
42	-0.047445	13263	sterite to .	0.418908	42	0.03022	1	!"
43	-0.021173	05919	*	0.419782	43	-0.08023	. **	
44	0.015799	0.04416	* *	0.419956	44	0.09951		l** .
45	0.021811	0.06097	I · I*	0.420053	45	-0.03340	. *	I .

Figure 4 Sample ACF and PACF of monthly logarithm transformation of American tourist visited for first seasonal differencing at period 6 months.

3.1.2 Model estimation and Evaluation

To check model adequacy, we consider whether the residuals of the model are white noise by using a Q statistic test with k = 12. The values of the Q statistic and p-values are given in Table 1. Procedure of choosing the models depends on the value of AIC and SBC (EKPENYONG, 2016). The model with the minimum values is considered as the best model for the data set (Scott, 2019). The models are presented in Table 2. From Table 2 the least AIC and SBC is ARIMA $(0, 0, 1) \times (1, 1, 0)$ 6 that indicates the ARIMA $(0, 0, 1) \times (1, 1, 0)$ 6 is the best model for forecasting the monthly American tourist visited during 2011 - 2015. (Yong & Brook, 2014).

Table 1The Q statistic test for k=12 of the tentative models for monthly American tourist visited.

Model	Q statistic	p-value
1. ARIMA $(1, 0, 0) \times (0, 1, 1)_6$	40.23	<.0001
2. ARIMA $(1, 0, 0) \times (0, 1, 2)_6$	31.25	0.0005
3. ARIMA $(4, 0, 0) \times (0, 1, 2)_6$	25.47	0.0006
4. ARIMA $(5, 0, 0) \times (0, 1, 1)_6$	47.51	<.0001
5. ARIMA $(5, 0, 0) \times (0, 1, 2)_6$	31.26	0.0005
6. ARIMA $(0, 0, 1) \times (1, 1, 1)_6$	12.51	0.1862
7. ARIMA $(0, 0, 1) \times (1, 1, 0)_6$	9.7	0.4676
8. ARIMA $(0, 0, 0) \times (1, 1, 0)_6$	14.77	0.1934
9. ARIMA $(0, 0, 0) \times (0, 1, 1)_6$	50.71	<.0001

Table 2 The summary of AIC and SBC for American visited.

Model	AIC	SBC
6. ARIMA $(0, 0, 1) \times (1, 1, 1)_6$	7.997593	13.6112
7. ARIMA $(0, 0, 1) \times (1, 1, 0)_6$	7.056717	10.79912
8. ARIMA $(0, 0, 0) \times (1, 1, 0)_6$	20.17794	22.04914

3.1.3 Diagnostic checking

In time series modeling the selection of best fit model is directly related to how well the residual analysis is performed. One of the assumptions of ARIMA model is that for a good model the residual should be white noise (Pelgrin, 2011).

Form the Figure 5 the sample ACF and PACF of the model shows that the autocorrelation of the residual are all close to zero which mean they are uncorrelated, hence the residual assume mean of zero and constant variance. Finally the p-value (0.4676) for the Ljung-Box statistic clearly exceeds 5% for all lag orders. Thus the selected model ARIMA $(0, 0, 1) \times (1, 1, 0)_6$ satisfies all the model assumptions.

Table 3 Parameter estimation of appropriate ARIMA $(0, 0, 1) \times (1, 1, 0)_6$

Model	MA1	SAR1
Parameter	0.87445	0.7458
SE	0.09300	0.10328

Autocorrelations

Partial Autocorrelations

Lag	Covariance	Correlation	-1 9 8 7 6 5 4 3 2 1 0 1 2 3 4 5 6 7	8 9 1 Std Error	Lag	Correlation	-198766	43210	1 2	3 4 5 6	5789	1
0	0.065113	1.00000	*********	*******								
1	-0.0061226	09403	. **	0.144338	1	-0.09403	I	. **				- 1
2	-0.0041889	06433	. *	0.146608	2	-0.07383	i	. *				i
3	-0.0014908	02290		0.146199	3	-0.03686	İ	. *				i
4	-0.0073639	11310	**	0.146274	4	-0.12589	i	. ***				i
5	0.0020888	0.03208		0.148084	5	0.00287	i	. i				i
6	0.00061626	0.00946		0.148229	6	-0.00606	i	- i				i
7	0.0011807	0.01813		0.148242	7	0.01343	i	. i				i
8	0.0046443	0.07133	i . i* .	0.148288	8	0.06455	i	. i	*			i
9	-0.0056242	08638	. **i .	0.149001	9	-0.06704	İ	. "				i
10	0.0053423	0.08205	i . i** .	0.150040	10	0.08193	i	. i	ilir Wr			i
11	0.015426	0.23692	, www.	0.150972	11	0.26223	i	. i	****			i
12	-0.014684	22551	.****	0.158529	12	-0.16876	i	. ***				i
13	0.014234	0.21861	www.	0.165077	13	0.23813	i	. i	****			i
14	-0.0013146	02019	i i i	0.171002	14	0.02730	i	. i	*			i
15	0.0016040	0.02463	i i i	0.171052	15	0.10478	i	. i	**			i
16	0.010112	0.15529	***	0.171126	16	0.17555	i	. i	***			i
17	-0.0012974	01993	i i i	0.174037	17	0.07662	i	. i	ir w			i
18	-0.0027117	04165	i *i .	0.174084	18	-0.02826	i	. *i				i
19	-0.0043378	06662	i *i .	0.174292	19	-0.04142	i	. *i				i
20	0.0064510	0.09907		0.174821	20	0.18589	i	. i	ir wwir			i
21	0.00098612	0.01514	i . i .	0.176987	21	-0.10097	i	. **				i
22	0.00012704	0.00195	1	0.176014	22	0.01026	i					i
23	-0.0017959	02758	. *1 .	0.176015	23	0.00944	i	. i				i
24	0.013992	0.21489	. ****	0.176105	24	0.09152	i	. i	ir-w			i
25	-0.0063493	09751	. ** .	0.181485	25	-0.00520	i	. i				i
26	0.00098022	0.01505	1	0.182574	26	-0.07168	i	. *i				i
27	0.0017838	0.02740	. * .	0.182600	27	-0.03078	i	. *i				i
28	-0.0037518	05762	. *1 .	0.182685	28	-0.06243	i	. *i				i
29	0.00011862	0.00182		0.183063	29	-0.02802	i	. *				i
30	0.00064478	0.00990		0.183064	30	-0.06158	i	. *i				i
31	0.0019305	0.02965	. * .	0.183075	31	-0.10252	i	. **				i
32	-0.0051418	07897	. ** .	0.183175	32	-0.08416	i	. **				i
33	0.0068204	0.10475	. ** .	0.183883	33	0.04757	i	. i	dr.			i
34	0.00053792	0.00826		0.186122	34	-0.00825	i					i
35	-0.0038921	05977	. *	0.186129	35	-0.19391	i	****				i
36	-0.0019339	02970	. *	0.186531	36	0.00850	i	. i				i
37	0.0062068	0.09532	**	0.186630	37	0.01658	i					i
38	0.0012701	0.01951		0.186647	38	0.04132	i	. i	*			i
39	0.0028212	0.04333		0.186690	39	0.09322	i		**			i
40	0.0017580	0.02700	*	0.186899	40	-0.01874	i					i
41	-0.0044530	06839		0.186980	41	-0.01542	i					i
42	-0.0075655	11619	. **	0.187501	42	-0.04015	i	. *1				i
43	-0.0040978	06293	. *	0.188995	43	0.01000	i					i
44	0.010306	0.16826		0.189431	44	0.02641	i	- i i				i
45	0.0018859	0.02896	I · I* ·	0.192165	45	0.14490	i		***			i
					-70	0						

Figure 5 Sample ACF and PACF of the residual of ARIMA $(0, 0, 1) \times (1, 1, 0)_6$

3.1.4 Forecasting

The forecast values with 95 percent forecast limit of the ARIMA $(0, 0, 1) \times (1, 1, 0)_{60}$ model for monthly American tourist are shown in Table 4 with standard error, lower and upper limit (Ekpenyong, 2016).

Table 4 Forecasted value for 6 months American tourist visited for 2015

Doto	Date Tarriet visited Foresasted value		95% confi	95% confidence limit		
Date	Tourist visited	Forecasted value	Lower	Upper	Error	
Jul-15	131	244.6	148.4	403.4	-113.6	
Aug-15	198	242.0	129.8	451.4	-44	
Sep-15	681	813.3	436.1	1386.2	-132.3	
Oct-15	1514	1449.4	777.1	2702.9	65.6	
Nov-15	1020	897.3	481.2	1673.4	122.7	
De-15	323	371.6	199.3	693.0	-48.6	

3.2 The best fitted model for each nationality

By considering all the steps in Box-Jenkins methodology the best fit model with t-ratio and variance of white noise forall the nationalities are shown in Table 5.

Table 5 The best fitted model for each nationality

Nationality/model	Parameters estimated $\hat{\sigma}_a^2$
1. American ARIMA (0, 0, 1) × (1, 1, 0) ₆	$(1+0.87445B^{6}) \ln X_{t} = (1+0.74358B)a_{t} 0.065113$ $(0.09300) \qquad (0.10328)$
2. British ARIMA (0, 0, 1) × (1, 1, 2) ₆	$(1+0.99004B^{6})\sqrt{X_{t}} = (1+0.39654B)(1-0.43849B^{6})a_{t}7.374285$ $(0.07764) \qquad (0.12700) \qquad (0.14988)$
3. Australian ARIMA (0, 0, 0) × (1, 1, 0) ₆	$(1+0.83277B^6) \ln X_t = a_t 0.120644$ (0.10493)
4. Singaporean ARIMA $(1, 1, 0) \times (1, 1, 0)_6$	$(1+0.54467B)(1+0.31214B^6)\sqrt{X_t} = a_t 6.467712$ $(0.11637) (0.12178)$
5. French ARIMA (0, 0, 0) × (1, 1, 0) ₆	$(1+0.94419B^6)\sqrt{X_t} = a_t 6.232704$ (0.07173)
6. German ARIMA $(0, 0, 0) \times (1, 1, 0)_6$	$(1+0.95822B^6)\sqrt{X_t} = a_t 19.81132$ (0.09604)
7. Chinese ARIMA (4, 1, 0) × (0, 0, 0)	$(1+0.50235B+0.29826B^2+0.81347B^3+0.65127B^4)X_t = a_t 24169.85$ $(0.12606) (0.11889) (0.11812) (0.15240)$
8. Thai ARIMA (1, 1, 1) × (0, 0, 0)	$(1+B)\sqrt{X_{t}} = (1+0.83645B)a_{t}0.692335$ $(0.03225) \qquad (0.11171)$
9. Malaysian ARIMA (5, 1, 1) × (0, 0, 0)	9.524783 $(1 + 0.23876B + 0.80638B^{2} + 0.70042B^{3} + 0.57061B^{4} + 0.43780B^{5})\sqrt{X_{t}} = (1 + 0.77508B)a_{t}$ $(0.17488) (0.22156) (0.23070) (0.21658) (0.14701)(0.17466)$
10. Japanese ARIMA (3, 1, 1) × (0, 0, 0)	$(1+0.38389B^{3})X_{t} = (1-0.52109B)a_{t} 31153.39$ $(0.13307) \qquad (0.11981)$
Total ARIMA $(0, 0,1) \times (1, 1, 0)_6$	$(1+0.79357B^{6})\sqrt{X_{t}} = 2.24357 - (1+0.47057B)a_{t}$ $(0.13374) \qquad (0.97511) \qquad (0.14447)$ 68.71606

4. Conclusion

In this study, a univariate time series models are selected by using the data of the past monthly tourist visited from top ten market sources during 2011 – 2015 obtained from Tourism Council of Bhutan. We applied Box-Jenkins model for forecasting the monthly tourist inflow in Bhutan. The graph, correlogram and periodogram of data show that some of nationalities data sets have seasonality at the period 6. From Table 5 we see that the best model for each nationality is divided into 2 groups. In the first group the model is ARIMA model which consist of Chinese -ARIMA $(4,1,0) \times (0,0,0)$, Thai - ARIMA $(1,1,1) \times (0,0,0)$, Malaysian - ARIMA $(5,1,1) \times (0,0,0)$ (0, 0, 0) and Japanese - ARIMA $(3, 1, 0) \times (0, 0, 0)$ and in the second group it is seasonal ARIMA model which consist of Australian-ARIMA $(0, 0, 0) \times (1, 1, 0)$ 6, British - ARIMA (0, 0, 0)1) \times (1, 1, 2)₆, American-ARIMA (0, 0, 1) \times (1, 1, 0)₆, Singaporean - ARIMA (0, 1, 1) \times (1, 1, $0)_6$, French - $(0, 0, 0) \times (1, 1, 0)_6$, German - ARIMA $(0, 0, 0) \times (1, 1, 0)_6$ and overall data -ARIMA $(0, 0, 1) \times (1, 1, 0)_{6}$. Among the visitors, Chinese, Thai, Malaysian and Japanese doesn't show seasonality. [Otieno, Mung'ta & Orwa, 2014). The main purpose of visitors from these four countries is to experience culture and tradition, and for spiritual and wellness activities which happen throughout the year so the season doesn't hinders the interest of their visit. But for American, Australian, British, Singaporean, French, and Germany shows seasonality as most of them preferred for adventurous tourism. The spring and autumn weather is basically warm and less rainfallwhich makes it favorable for the visitors those who are interested in (Tourism council of Bhutan, n.d.).

5. Acknowledgements

We would like to thank Thai International Cooperation Agency (TICA) for providing the financial support. We also would like to thank Tourism Council of Bhutan (TCB) for providing us latest available data.

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